

Modeling of Data

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- *NRiC* §15.
- Model depends on adjustable parameters.
- Can be used for “constrained interpolation.”
- Basic approach:
 1. Choose *figure-of-merit* function (e.g., χ^2).
 2. Adjust *best-fit parameters*: minimize merit function.
 3. Compute *goodness-of-fit*.
 4. Compute *error estimates* for parameters.

Least Squares Fitting

- Suppose we want to fit N data points (x_i, y_i) with a function that depends on M parameters a_j and that each data point has a standard deviation σ_i . The *maximum likelihood estimate* of the model parameters is obtained by minimizing:

$$\chi^2 \equiv \sum_{i=1}^N \left[\frac{y_i - y(x_i; a_1 \dots a_M)}{\sigma_i} \right]^2.$$

- Assuming the errors are normally distributed, a “good fit” has $\chi^2 \sim \nu$, where $\nu = N - M$.
 - NOTE: Assumption of normal errors means glitches or outliers in data may overbias the fit—see *NRiC* §15.7 for discussion of more robust methods.
 - Grossly overestimated (underestimated) σ_i 's may give incorrect impression that fit is very good (very bad).

- If uncertain about reliability of goodness-of-fit measure, could do *Monte Carlo simulations* of fits to synthetic data.
- Question: what to do if σ_i 's not known? Answer: choose an arbitrary constant σ , perform the fit, then estimate σ from the fit:
$$\sigma^2 = \sum_{i=1}^N [y_i - y(x_i)]^2 / \nu$$
 (note the denominator is what χ^2 *should* approximately be equal to, if the fit is good).

Fitting Data to a Straight Line (Linear Regression)

- For this case the model is simply:

$$y(x) = y(x; a, b) = a + bx,$$

and

$$\chi^2(a, b) = \sum_{i=1}^N \left(\frac{y_i - a - bx_i}{\sigma_i} \right)^2.$$

- Derive formula for best-fit parameters by setting $\partial\chi^2/\partial a = 0 = \partial\chi^2/\partial b$. See *NRiC* §15.2 for the derivation (note: `sm` uses the same formulae for its `lsq` routine).

- Derive uncertainties in a and b from propagation of errors:

$$\sigma_f^2 = \sum_{i=1}^N \sigma_i^2 \left(\frac{\partial f}{\partial y_i} \right)^2,$$

where $f = a(x_i, y_i, \sigma_i), b(x_i, y_i, \sigma_i)$ in this case (the x_i 's have no uncertainties).

- Want probability that χ^2 is bad by chance

$$Q = \text{gammq}((N - 2)/2, \chi^2/2) > 10^{-3} \text{ (here } (N - 2)/2 \equiv \nu/2\text{)}.$$

General Linear Least Squares

- Can generalize to any combination that is linear in a_j 's:

$$y(x) = \sum_{j=1}^M a_j X_j(x),$$

e.g., $y(x) = a_1 + a_2x + a_3x^2 + \dots + a_Mx^{M-1}$, or sines and cosines.

- Define $N \times M$ *design matrix* $A_{ij} = X_j(x_i)/\sigma_i$. Note $N \geq M$ for the fit to make sense.
- Also define vector \mathbf{b} of length N where $b_i = y_i/\sigma_i$, and vector \mathbf{a} of length M where $a_i = a_1, \dots, a_M$.
- Then we wish to find \mathbf{a} that minimizes:

$$\chi^2 = |\mathbf{Aa} - \mathbf{b}|^2.$$

- This is what SVD solves!

- Recall for SVD we had $\mathbf{A} = \mathbf{U}\mathbf{W}\mathbf{V}^T$.
- Rewriting the SVD solution we get:

$$\mathbf{a} = \sum_{j=1}^M \left(\frac{\mathbf{U}_{(j)} \cdot \mathbf{b}}{w_j} \right) \mathbf{V}_{(j)},$$

where $\mathbf{U}_{(j)}$ (length N) and $\mathbf{V}_{(j)}$ (length M) denote columns of \mathbf{U} and \mathbf{V} , respectively.

- As before, if w_j is small (or zero), can set $1/w_j = 0$.
 - Useful because least-squares problems are generally *both* overdetermined ($N > M$) *and* underdetermined (ambiguous combinations of parameters exist)!
- Can also compute variances of estimated parameters:

$$\sigma^2(a_j) = \sum_{i=1}^M (V_{ji}/w_i)^2.$$
- Can generalize to multidimensions.

Nonlinear Models

- Suppose model depends *nonlinearly* on the a_j 's, e.g.,
 $y(x) = a_1 \sin(a_2 x + a_3)$.
- Still minimize χ^2 , but must proceed iteratively:
 - Use $\mathbf{a}_{\text{next}} = \mathbf{a}_{\text{cur}} - \lambda \nabla \chi^2(\mathbf{a}_{\text{cur}})$ far from minimum (steepest descent), where λ is a constant.
 - Use $\mathbf{a}_{\text{next}} = \mathbf{a}_{\text{cur}} - \mathbf{D}^{-1}[\nabla \chi^2(\mathbf{a}_{\text{cur}})]$ close to minimum, where \mathbf{D} is the *Hessian* matrix.

- ● D comes from considering Taylor series expansion of $f(\mathbf{x})$ near a point \mathbf{P} :

$$\begin{aligned} f(\mathbf{x}) &= f(\mathbf{P}) + \sum_i \frac{\partial f}{\partial x_i} x_i + \frac{1}{2} \sum_{i,j} \frac{\partial^2 f}{\partial x_i \partial x_j} x_i x_j + \dots \\ &\simeq c - \mathbf{b} \cdot \mathbf{x} + \frac{1}{2} \mathbf{x} \mathbf{A} \mathbf{x}, \end{aligned}$$

where $c \equiv f(\mathbf{P})$, $\mathbf{b} \equiv -\nabla f|_{\mathbf{P}}$, and $A_{ij} \equiv \left. \frac{\partial^2 f}{\partial x_i \partial x_j} \right|_{\mathbf{P}}$. Here \mathbf{A} is the Hessian matrix. Note that $\nabla f = \mathbf{A}\mathbf{x} - \mathbf{b}$.

- Close to its minimum, χ^2 can be approximated by the above quadratic form, and so an “exact” step can be taken to get to the point where $\nabla\chi^2 = 0$. The step is just $\mathbf{x}' - \mathbf{x} = -\mathbf{A}^{-1}\nabla f|_{\mathbf{P}}$.
- In practice, terms involving the second derivatives of y with respect to the fit parameters can be ignored, so the Hessian matrix is much simpler to compute (recall the χ^2 function contains the model y).
- The *Levenberg-Marquardt method* adjusts λ to smooth the transition between these two regimes (vary between a diagonal matrix and inverse Hessian).
 - Cf. *NRiC* §15.5 for details of the L-M method.

Levenberg-Marquardt method in NRiC

- *NRiC* provides two routines, `mrqmin()` and `mrqcof()`, that implement the L-M method.
- The user must provide a function that computes $y(x_i)$ as well as all the partial derivatives $\partial y / \partial a_j$ evaluated at x_i .
- The routine `mrqmin()` is called iteratively until a successful step (i.e., one in which λ gets smaller) changes χ^2 by less than a fractional amount, like 0.001 (no point in doing better).

- Points to consider:
 - The argument list for `mrqmin()` is *very* complicated. For example, you can request that some parameters be held fixed (via input array `ia`).
 - You need to specify an initial guess for each a_j (and set $\lambda < 0$).
 - Estimated variances in the parameters are returned as the diagonal elements of the *covariance matrix* (`covar`), if you call `mrqmin()` with $\lambda = 0$.
 - Also calls *NRiC* routines `covsrt()` and `gaussj()`.

```

void mrqmin(float x[], float y[], float sig[], int ndata, float a[], int ia[],
           int ma, float **covar, float **alpha, float *chisq,
           void (*funcs)(float, float [], float *, float [], int), float *alamda)
/* Levenberg-Marquardt method, attempting to reduce the value of Chi^2 of a fit
between a set of data points x[1..ndata], y[1..ndata] with individual standard
deviations sig[1..ndata], and a nonlinear function dependent on ma coefficients
a[1..ma]. The input array ia[1..ma] indicates by nonzero entries those
components of a that should be fitted for, and by zero entries those components
that should be held fixed at their input values. The program returns current
best fit values of the parameters a[1..ma], and Chi^2=chisq. ...
Supply a routine funcs(x,a,yfit,dyda,ma) that evaluates the fitting function
yfit, and its derivatives dyda[1..ma] with respect to the fitting parameters a
at x. On the first call provide an initial guess for the parameters a, and set
alambda<0 for initialization (which sets alambda=0.001). If a step succeeds
chisq becomes smaller and alambda decreases by a factor of 10. If a step fails
alambda grows by a factor of 10. You must call this routine repeatedly until
convergence is achieved. Then, make a final call with alambda=0, so that
covar[1..ma][1..ma] returns the covariance matrix, and alpha the curvature
matrix. (Parameters held fixed will return zero covariances.) */
{
    void covsrt(float **covar, int ma, int ia[], int mfit);
    void gaussj(float **a, int n, float **b, int m);
    void mrqcof(float x[], float y[], float sig[], int ndata, float a[],
               int ia[], int ma, float **alpha, float beta[], float *chisq,
               void (*funcs)(float, float [], float *, float [], int));
    .....
    .....
void fgauss(float x, float a[], float *y, float dyda[], int na)
//The dimensions of the arrays are a[1..na], dyda[1..na].
{
    .....
    .....
}

```